

# Binomial Line Cox Processes

This talk: Part 1 - Characterization and Meta-Distributions

Gourab Ghatak

Department of Electrical Engineering  
IIT Delhi

June 4, 2024

Part 2 - *Applications in Wireless Networks and Automotive Radars* is a separate talk.

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## Acknowledgements

- Joint work with Md. Taha Shah (Ph.D. student at IIT Delhi) and Martin Haenggi (EE, University of Notre Dame).
- Thanks to Maruti-Suzuki India Limited for supporting a part of this work.
- The talk (Part 1) is derived from the following three papers:
  - G. Ghatak, *Binomial Line Processes: Distance Distributions*, IEEE Transactions on Vehicular Technology, doi: 10.1109/TVT.2021.3134834.
  - M. T. Shah, et al. *Analyzing Wireless Networks using Binomial Line Cox Processes* in the IEEE WiOpt 2023 Workshop on Spatial Stochastic Models for Wireless Networks - SpaSWiN.
  - Md. T. Shah, et al. *Binomial Line Cox Processes: Statistical Characterization and Applications in Wireless Network Analysis*, IEEE Transactions on Wireless Communication (accepted, May 2024).

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## Overview of Today's Talk

1. Context
2. Motivation and History of Line Processes
3. Binomial Line Processes
4. Binomial Line Cox Process
5. Meta Distribution of the SINR with BLCP Nodes

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## Context

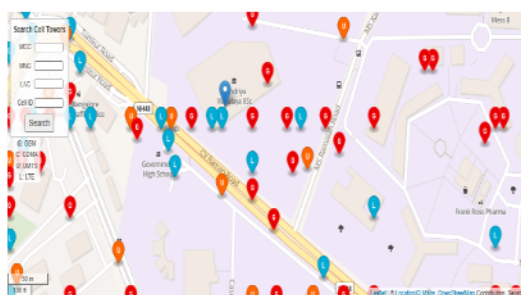
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### Modern Networks are Increasingly AdHoc



(a)



(b)

Figure: BS discovery at (a) IISc Campus and (b) ISI Bangalore Campus. Source: OpenCellID.

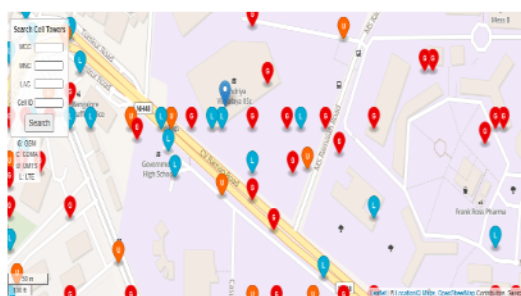
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### Modern Networks are Increasingly AdHoc



(a)



(b)

Figure: BS discovery at (a) IISc Campus and (b) ISI Bangalore Campus. Source: OpenCellID.

- Are these two scenarios to be studied separately?

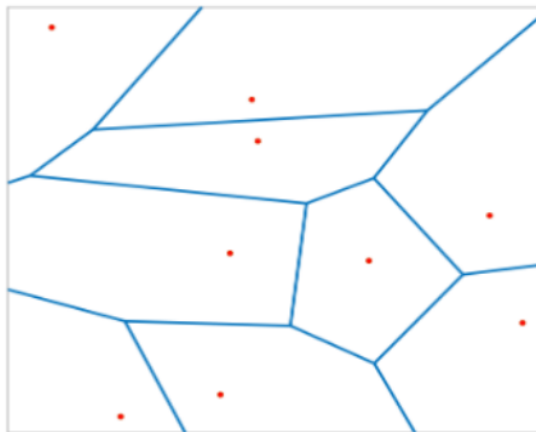
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## Solution - Stochastic Networks and Point Processes



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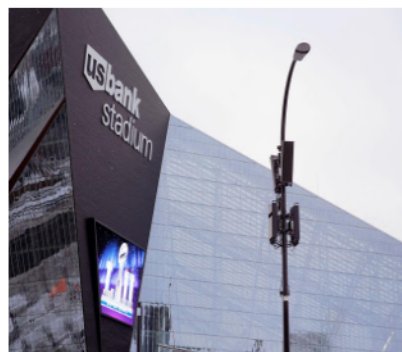
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## Next-Generation Wireless Deployment



(a)



(b)

Figure: (a) Example first-generation mm-Wave deployment. (b) Verizon's small cell deployment near stadiums.

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## Other examples

- New York City has already deployed such small cells focusing on high-speed connectivity for pedestrian users.
- South Korea: SK Telecom, 28GHz, outdoor pedestrian.
- Taiwan: Nokia and CHT.
- Verizon deployed 5G (with mm-wave) on street lights in Sacramento.
- AT&T deployed 5G (with mm-wave) on smart lamp-posts in San Jose.

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Ok, so deployment along streets. What else?

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## Coverage Footprint from a 5G Lamppost

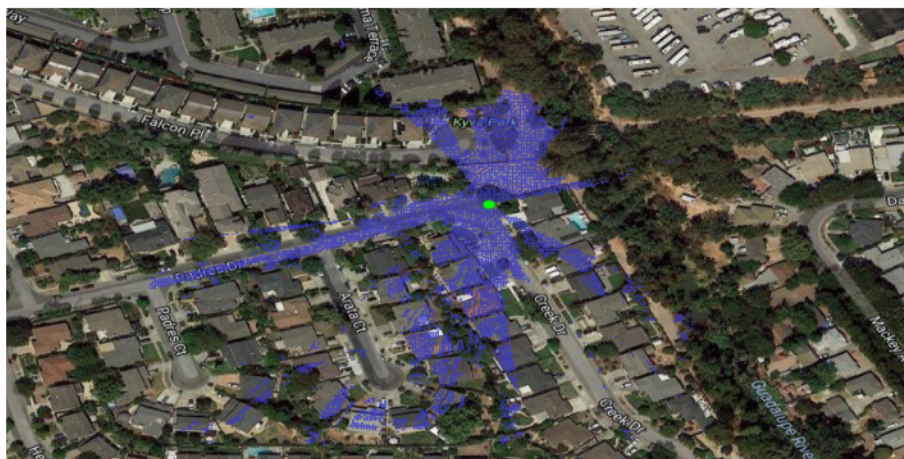


Figure: Coverage of a lamp-post mm-wave small cell in San Jose.

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## Motivation and History of Line Processes

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## Networks on Lines - Other Examples

- Movement of vehicular nodes.

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<sup>1</sup>Will be made rigorous in a moment.

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## Networks on Lines - Other Examples

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- Autonomous robots movements:
  - Inspecting underground pipelines.
  - Movement in warehouses.

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Can be modeled, studied, and characterized using **line processes**<sup>1</sup> (one candidate): a random collection of lines in, e.g., Euclidean space (e.g., a plane).

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Let's look at some history...

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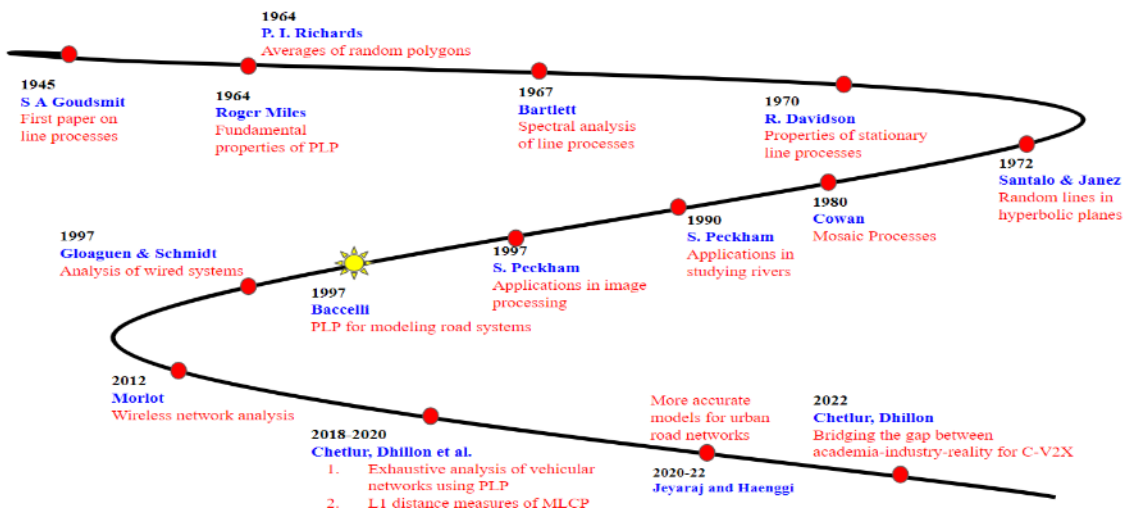
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# History of Line Process Research



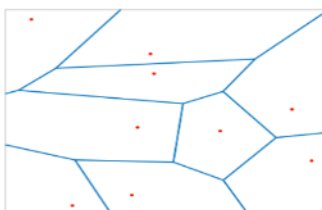
[A part of the illustration is from Prof. Dhillon's book/talk.]

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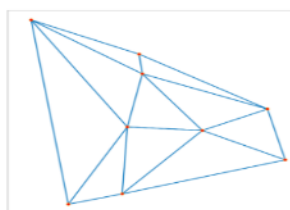
# Candidate Models for Emulating Streets

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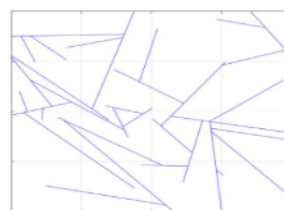
# Candidate Models for Emulating Streets



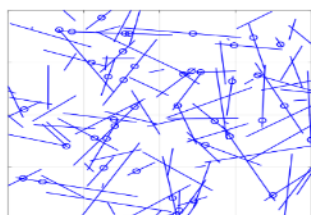
Poisson Voronoi Tessellation



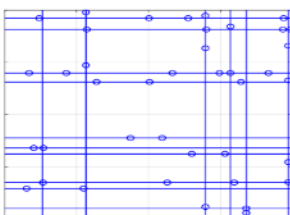
Poisson Delaunay Triangulation



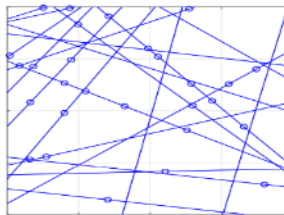
Poisson Lilypond Model



Poisson Stick Process



Manhattan Line Process



Poisson Line Process

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## Some Observations

- Homogeneity → characterization is motion-invariant - one can study the network from the perspective of a single point, e.g., at the origin.

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## Some Observations

- Homogeneity → characterization is motion-invariant - one can study the network from the perspective of a single point, e.g., at the origin.
- From the perspective of a single city, the street networks are **inhomogeneous** - denser streets in the city center, not so much in suburbs.

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## Some Observations

- Homogeneity → characterization is motion-invariant - one can study the network from the perspective of a single point, e.g., at the origin.
- From the perspective of a single city, the street networks are **inhomogeneous** - denser streets in the city center, not so much in suburbs.
- Solutions need to be adaptive -
  - MAC protocols should adapt based on the location.
  - Deployment planning must take into account the location - e.g., where is the nearest eV charging point? nearest bus stop?
  - V2X load is directly dependent on the length of streets in a coverage area; load-balancing and cell-breathing need to be adaptive.
  - ...

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# Emulating Non-Homogeneous Streets

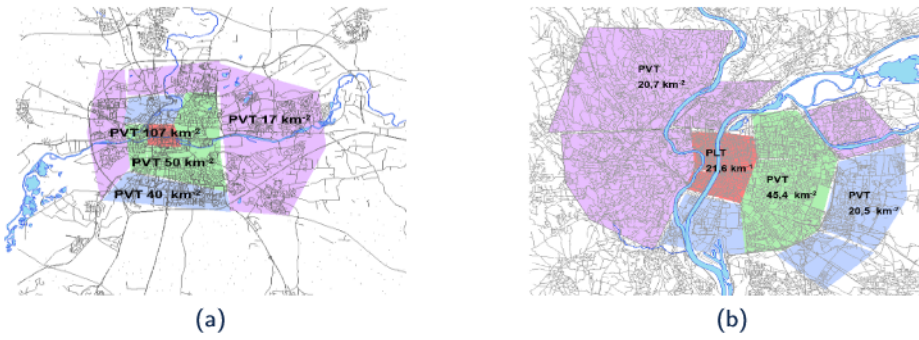
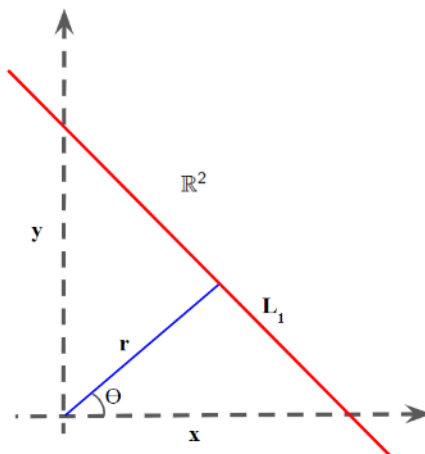


Figure: (a) PVT and (b) PVT + PLT models for fitting for streets of Lyon <sup>2</sup>. For a PVT, the model parameter is the density of the underlying Poisson point process (PPP).

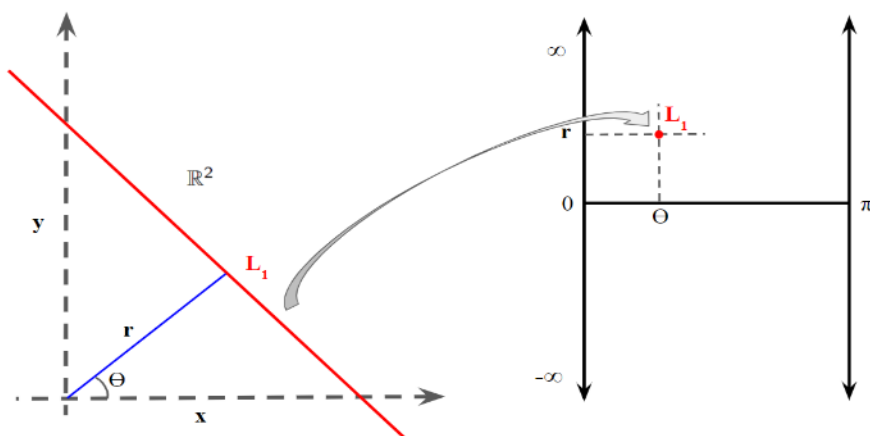
<sup>2</sup>C. Gloaguen, "Modélisation macroscopique géométrique des réseaux d'accès en télécommunication," Presentation at Société de Mathématiques Appliquées et Industrielles. Link: <http://smai.emath.fr/spip/IMG/ppt>, 2010

# Binomial Line Processes

# Line Process - Construction

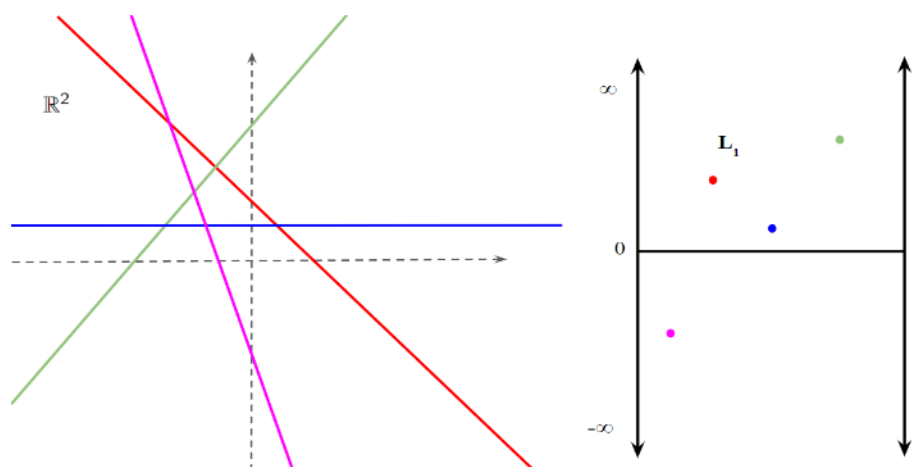


## Line Process - Construction



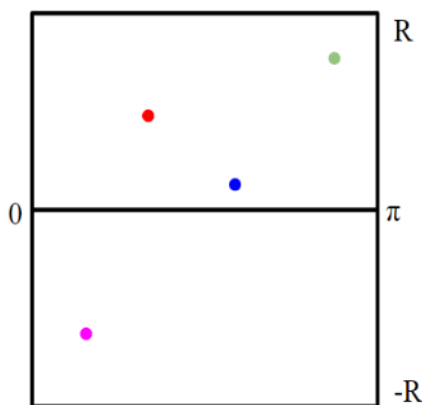
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## Line Process - Construction



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## Binomial Line Process - Construction



- **Key:** Fixed number of points ( $n_B$ ), uniformly distributed in a closed compact set  $\mathcal{D} \triangleq [-R, R] \times [0, \pi)$ .

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## BLP - Definition

### Binomial Line Process (BLP)

A binomial line process (BLP)  $\mathcal{L}$  is a collection of  $n_B$  lines in the two-dimensional Euclidean plane. Formally,

$$\mathcal{L} \subset \mathcal{Q} \triangleq \bigcup_{r \in [-R, R], \theta \in [0, \pi)} \{(x, y) \in \mathbb{R}^2 : x \cos \theta + y \sin \theta = r\}. \quad (1)$$

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- We will call  $\mathcal{D}$  the generating set or the domain set of  $\mathcal{L}$ , and

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- A point  $(\theta_i, r_i) \in \mathcal{D}$ , corresponding to a line  $L_i \in \mathcal{L}$ , the generating point of  $L_i$ .

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- Generating points form a BPP in  $\mathcal{D}$ .

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## Instance Visualization

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## Instance Visualization

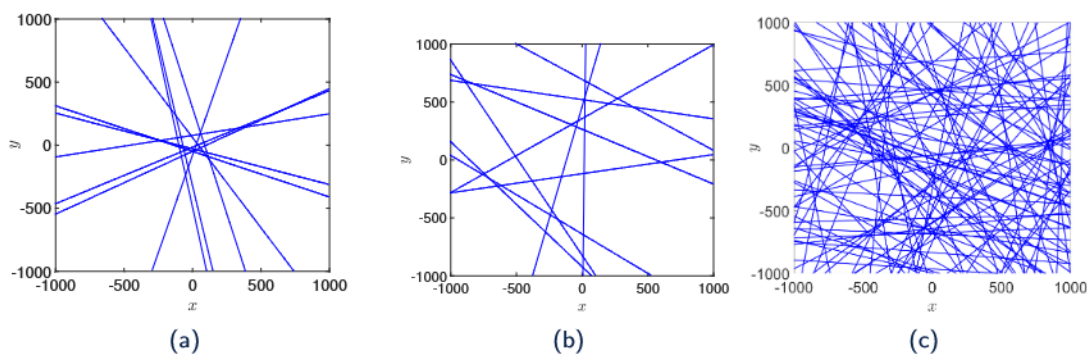


Figure: (a) A BLP with  $n_B = 10$  and  $R = 100$ , (b) A realization of a PLP conditioned on 10 lines generated in the window of interest, and (c) A PLP whose underlying PPP has intensity  $\frac{n_B}{2\pi R}$  with  $n_B = 10$  and  $R = 100$ . Note: Here  $R$  is the radius of the circle in which BLP lines are generated.

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## Non-stationarity

- BLP is a non-stationary process.

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- BLP is a non-stationary process.
- $\implies$  the statistics of the BLP cannot be characterized from the perspective of a single typical point located, say, at the origin.

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# Non-stationarity

- BLP is a non-stationary process.
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- However, isotropic  $\rightarrow$  the properties as seen from a point depends only on its distance from the origin and not its orientation.
- Accordingly, without loss of generality, let us consider a test point located at  $(0, r_0)$ .

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## Distance to the Nearest BLP Line

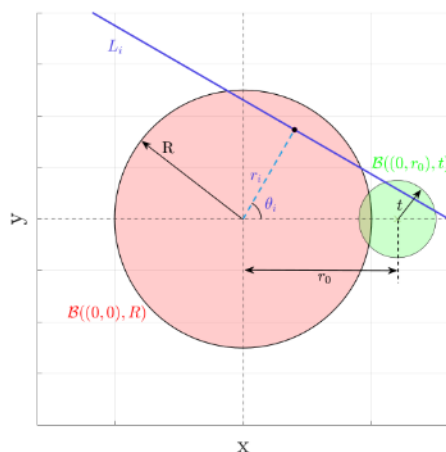


Figure: Illustration of the construction of a BLP and intersecting lines on  $B((0, r_0), t)$ .

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## Domain Bands

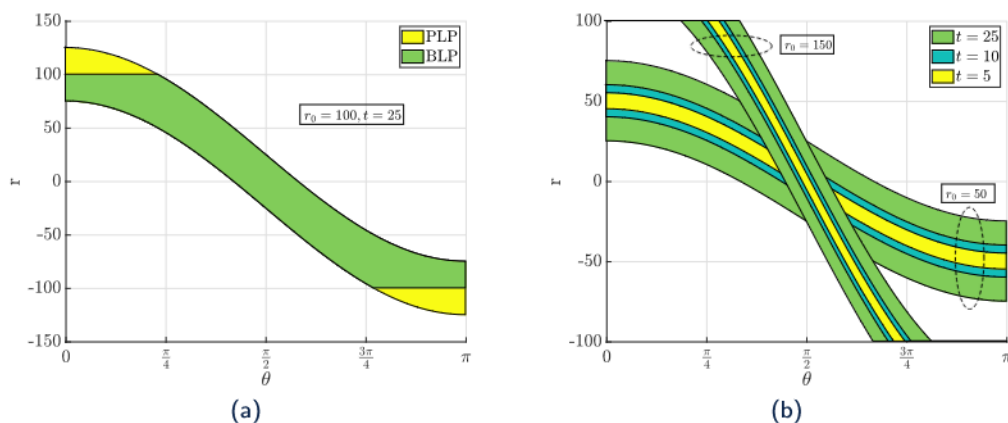


Figure: (a) Illustration of the domain bands for a PLP and a BLP with  $R = 100$ . (b) Domain bands for different values of  $t$  and  $r_0$ . Here  $R = 100$ . Note that when  $r_0 + t \leq R$ , the domain bands for PLP and BLP coincide.

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## Area of the Domain Band

### Theorem

The area of the domain band  $\mathcal{D}_B(r_0, t)$  for a BLP is

$$A_D(r_0, t) = \begin{cases} 2\pi t; & \text{for } r_0 + t \leq R \\ 2\pi t - 2r_0 \sqrt{1 - \left(\frac{R-t}{r_0}\right)^2} + 2(R-t) \cos^{-1}\left(\frac{R-t}{r_0}\right); & \text{for } r_0 + t > R \text{ and } r_0 - t \leq R \\ 2\pi t - 2r_0 \left( \sqrt{1 - \left(\frac{R-t}{r_0}\right)^2} - \sqrt{1 - \left(\frac{R+t}{r_0}\right)^2} \right) + 2(R-t) \cos^{-1}\left(\frac{R-t}{r_0}\right) \\ \quad - 2(R+t) \cos^{-1}\left(\frac{R+t}{r_0}\right); & \text{for } r_0 - t \geq R. \end{cases} \quad (2)$$

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## Distance to the nearest street?

### Corollary

[Void Probability] The probability that no line of the BLP intersects with  $\mathcal{B}((0, r_0), t)$  is

$$\mathcal{V}_{\text{BLP}}(n_B, \mathcal{B}((0, r_0), t)) = \left(1 - \frac{A_D(r_0, t)}{2\pi R}\right)^{n_B},$$

where,  $n_B$  is the number of lines of the BLP  $\mathcal{L}$  and  $A_D$  is area of the domain bands.

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### Corollary

The CDF of the distance to the nearest line of the BLP from a test point at  $(0, r_0)$  is,

$$F_d(t) = 1 - \mathcal{V}_{\text{BLP}}(n_B, \mathcal{B}((0, r_0), t)).$$

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## Line Length Density

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### Definition

The line length measure is

$$\mathcal{R}(S) = n_B \mathbb{E}(|L \cap S|_1), \quad S \subset \mathbb{R}^2,$$

where  $|\cdot|_1$  is the Lebesgue measure in one dimension and  $L$  is a line of the BLP. The corresponding radial density is

$$\rho(r) = \lim_{u \rightarrow 0} \frac{\mathcal{R}(\mathcal{B}((0,0), r+u) \setminus \mathcal{B}((0,0), r))}{\pi(2u + u^2)}.$$

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The line length measure follows by integrating  $\rho(r)$ , i.e.,

$$\mathcal{R}(S) = \int_S \rho(|\mathbf{x}|) \, d\mathbf{x}, \quad S \subset \mathbb{R}^2.$$

## Line Length Density

### Theorem

For a BLP generated by  $n_B$  lines within a disk of radius  $R$ ,

$$\rho(r) = \begin{cases} \frac{n_B}{2R}, & \text{if } r \leq R \\ \frac{n_B}{\pi R} \arcsin\left(\frac{R}{r}\right) & \text{if } r > R. \end{cases}$$

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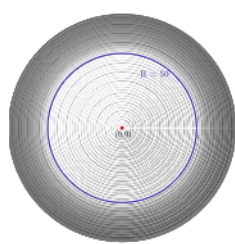
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## Line Length Density

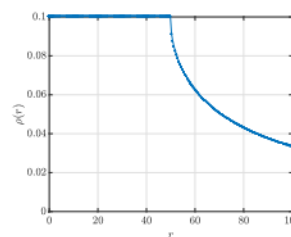
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(a)



(b)

Figure: (a) Ratio of the line length measure to the area in concentric annuli of equal width  $w = 2$ . Here,  $R = 50$ . (b) Line length density  $\rho(r)$  for  $R = 50$  and  $n_B = 10$ .

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## Intersection Density

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## Intersection Density

### Theorem

The intersection density for a PLP with density  $\lambda_{PPP}$  is

$$\rho_P(\lambda_{PPP}) = \pi \lambda_{PPP}^2.$$

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## Intersection Density

### Theorem

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### Theorem

The radial intersection density at a distance  $r$  from the origin for a BLP generated by  $n_B$  lines within a disk of radius  $R$  is

$$\rho_{\times}(r) = \begin{cases} \frac{n_B(n_B-1)}{4\pi R^2}, & \text{if } r \leq R, \\ \frac{n_B(n_B-1)}{4\pi^2 R^2 r} \left( 2r \arcsin\left(\frac{R}{r}\right) - \frac{2R}{r} \sqrt{r^2 - R^2} \right) & \text{if } r > R. \end{cases}$$

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## Radial Intersection Density: Proof

- Let  $S = \mathcal{B}((0,0), t)$ , and consider  $L_0$  generated at  $(0, r_0)$  where  $0 \leq r_0 \leq \min\{t, R\}$ .
- First, we determine the domain band  $\mathcal{D}_{\times}$  corresponding to the intersection on  $L_0$
- For a  $\theta$ , the range of  $r$  that correspond to the domain band is,
 
$$\max \left\{ -R, \left( r_0 \cos \theta - \sqrt{t^2 - r_0^2 \sin^2 \theta} \right) \right\} \leq r_i \leq \min \left\{ R, \left( r_0 \cos \theta + \sqrt{t^2 - r_0^2 \sin^2 \theta} \right) \right\}.$$
- For  $t > R$ , the domain band gets clipped to  $R$  (upper) and  $-R$  (lower) limits

$$A_{D_{\times}}(t) = \begin{cases} \pi t, & \text{if } t \leq R, \\ \frac{2}{R} \left( t^2 \arcsin\left(\frac{R}{t}\right) + 2R^2 \arccos\left(\frac{R}{t}\right) - \sqrt{t^2 - R^2} \right) & \text{if } t > R. \end{cases}$$

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## Radial Intersection Density: Proof

- Accordingly, the probability that a line intersects  $L_0$  within  $S$  is obtained as

$$\mathcal{P}_\times(t) = \frac{A_{D_\times}(t)}{2\pi \min\{t, R\}}.$$

- Now, let us assume that  $k$  lines are generated in  $S$ . Each of these intersects  $L_0$  with probability  $\mathcal{P}_\times(t)$ . As a result, the average number of intersections on  $L_0$  within  $S$  from the  $k$  lines is

$$\mathcal{N}' = \sum_{j=0}^k j \binom{k}{j} (\mathcal{P}_\times(t | t \leq R))^j (1 - \mathcal{P}_\times(t | t \leq R))^{k-j} = \frac{k}{2}.$$

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## Radial Intersection Density: Proof

Finally, in order to determine the average number of intersections on all the lines within  $S$ , we take the expectation over the number of lines generated within  $S$ . This is evaluated as

$$\mathcal{N}_1 = \sum_{k=0}^{n_B-1} \underbrace{\binom{n_B}{k+1}}_{T_1} \underbrace{\left(\frac{t}{R}\right)^{k+1} \left(1 - \frac{t}{R}\right)^{n_B-k-1}}_{T_2} \underbrace{\frac{k}{2}}_{T_3} \times \underbrace{(k+1)}_{T_4} \times \underbrace{\frac{1}{2}}_{T_5} = \frac{n_B(n_B-1)}{4} \left(\frac{t}{R}\right)^2. \quad (3)$$

Similarly, for  $t > R$

$$\begin{aligned} \mathcal{N}_2 &= \sum_{k_1=0}^{n_B-1} k_1 \mathcal{P}_\times(t | t > R) \\ &= \frac{n_B(n_B-1)}{2\pi R^2} \times \left( t^2 \arcsin\left(\frac{R}{t}\right) + 2R^2 \arccos\left(\frac{R}{t}\right) - R\sqrt{t^2 - R^2} \right). \end{aligned} \quad (4)$$

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## Intersection Density

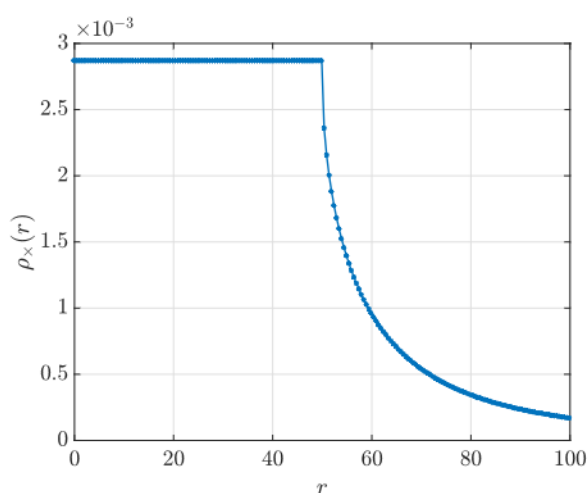


Figure: Intersection density for  $R = 50$  and  $n_B = 10$ .

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## Highlights of the Properties

- The line length density as well as the density of intersections remain constant at  $\frac{n_B}{2R}$  for  $r \leq R$  and then decreases as  $\mathcal{O}(1/r)$  as  $r \rightarrow \infty$ .

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- The line length density as well as the density of intersections remain constant at  $\frac{n_B}{2R}$  for  $r \leq R$  and then decreases as  $\mathcal{O}(1/r)$  as  $r \rightarrow \infty$ .
- Other results skipped: distance to the nearest intersection.

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## Binomial Line Cox Process

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## BLCP - Construction

- On each  $L_i$  of  $\mathcal{L}$ , define an independent 1D PPP  $\Phi_i$  with intensity  $\lambda$ .

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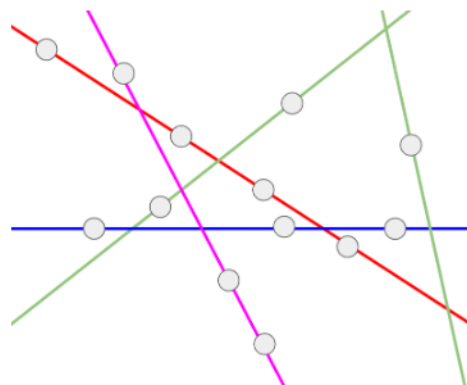
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## Nearest BLCP Point - Chord Length

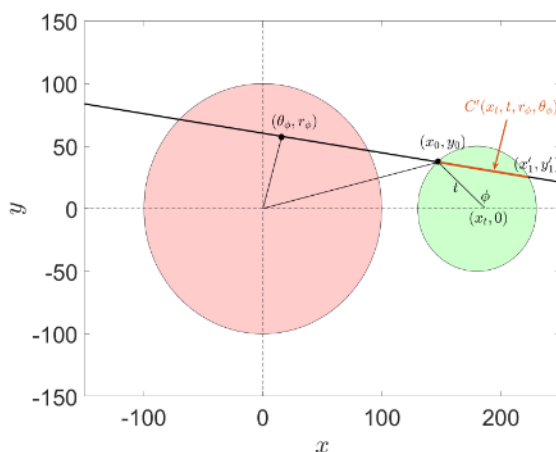


Figure: Illustration of the chord length  $C'(x_t, t, r_\phi, \theta_\phi)$ .

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## Void Probability

### Theorem

The probability that the disk  $\mathcal{B}((0, r_0), t)$  contains no points of  $\Phi$  is given by

$$\mathcal{V}_{BLCP}(n_B, \mathcal{B}((0, r_0), t)) = \left[ \frac{1}{2\pi R} \iint_{\mathcal{D}(r_0, t)} \underbrace{\exp(-\lambda C(\theta, r))}_{\mathcal{V}_{PPP}(\text{Chord Length})} dr d\theta \right]^{n_B}, \quad (5)$$

where,

$$C(\theta, r) = \begin{cases} 2\sqrt{t^2 - (r_0 \cos \theta - r)^2}; & t \geq |r_0 \cos \theta - r|, \\ 0; & \text{otherwise,} \end{cases} \quad (6)$$

is the length of the chord created by a line corresponding to  $(\theta, r) \in \mathcal{D}$  in the disk  $\mathcal{B}((0, r_0), t)$ .

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## Void Probability

### Corollary

*[Distance Distribution] Following the void probability, the distance distribution of the nearest BLCP point from the test point  $(0, r_0)$  is*

$$F_{d_1}(t) = 1 - \mathcal{V}_{BLCP}(n_B, \mathcal{B}((0, r_0), t)).$$

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From a wireless network perspective where the locations of the APs are modeled as points of a BLCP, the above result characterizes the distance distribution to the nearest AP.

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## Palm Perspective of the BLCP

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## Palm Perspective of the BLCP

### Lemma

For a BLCP  $\Phi$  defined on a BLP  $\mathcal{P}$  with  $n_B$  lines, we have

$$\mathbb{P}(\Phi \in \mathcal{Y} \mid \mathbf{x} \in \Phi) = \mathbb{P}(\Phi_{n_B-1} \cup \Phi_{\mathbf{x}} \cup \{\mathbf{x}\} \in \mathcal{Y}), \quad (7)$$

where  $\Phi_{\mathbf{x}}$  is a 1D PPP on a randomly oriented line that passes through  $\mathbf{x}$ .

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Necessary to characterize the statistics of the performance metrics conditioned on an event that a node, e.g., a transmitter or an AP is located at a given point.

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Necessary to characterize the statistics of the performance metrics conditioned on an event that a node, e.g., a transmitter or an AP is located at a given point.

- Consider a network - APs  $\sim$  BLCP.
- If a Rx at the origin associates with a Tx located at  $\mathbf{x}$ , then the interfering APs are located not only in the other  $n_B - 1$  lines but also on a line necessarily passing through  $\mathbf{x}$ .

## Probability Generating Functional - Definitions

### Definition

The **PGFL** of a point process  $\Phi$  evaluated for a function  $\nu$  is defined mathematically as the Laplace functional of  $-\log \nu$  and is calculated as  $\mathbb{E} [\prod_{\mathbf{x} \in \Phi} \nu(\mathbf{x})]$ .

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Characterizes the behavior of any random point function

- Given pollution sources and diffusion models, the expected intensity of pollution at a location.
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For a PPP

$$\mathbb{E} \left[ \prod_{\mathbf{x} \in \Phi_{PPP}} \nu(\mathbf{x}) \right] = \exp \left( - \int_{\mathbb{R}^2} 1 - \nu(\mathbf{x}) \Lambda(d\mathbf{x}) \right) \quad (8)$$

## Probability Generating Functional

For a particular  $r_0$  and  $d_1$ ,

- a line is intersecting if  $|r_0 \cos \theta - r| \geq d_1$  and non-intersecting otherwise.

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$$G_I(r_0, d_1) = \frac{1}{A_D(r_0, d_1)} \iint_{\mathcal{D}_B(0, d_1)} \exp \left( -2\lambda \int_{\sqrt{d_1^2 - (r_0 \cos \theta - r)^2}}^{\infty} 1 - f \left( \sqrt{y^2 + (r_0 \cos \theta - r)^2} \right) dy \right) dr d\theta,$$

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$$G_{NI}(r_0, d_1) = \frac{1}{(2\pi R - A_D(r_0, d_1))} \iint_{\mathcal{D} \setminus \mathcal{D}_B(0, d_1)} \exp \left( -2\lambda \int_0^{\infty} 1 - f \left( \sqrt{y^2 + (r_0 \cos \theta - r)^2} \right) dy \right) dr d\theta. \quad (9)$$

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The line containing the nearest point intersects the disk  $\mathcal{B}((0, 0), d_1)$  almost surely. Whereas, the other  $n_B - 1$  lines may or may not intersect. Accordingly,

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## What could we characterize

- Distance to the nearest street as we move away from the city center.

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Let's look at an application with meta distributions.

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## Meta Distribution of the SINR with BLCP Nodes

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### Refresher On Indicators and Conditional Expectations

- For a random variable,  $Z$ ,  $\mathbb{E}[Z]$  is the mean of  $Z$ .

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  - The probability of any event can be expressed by adding or subtracting these elementary probabilities.
- However, if  $Z$  is a function of other sources of randomness, then  $\mathbb{E}[\mathbb{1}(Z > z)]$  does not reveal how the statistics of  $Z$  depend on those of the individual random elements.

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## Meta-Distributions

- Let us take the example of  $Z = f(X, Y)$  where  $X$  and  $Y$  are independent.

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$$\bar{F}_{\llbracket Z|Y \rrbracket}(z, x) = \mathbb{E} [\mathbb{1} (\mathbb{E} [\mathbb{1} (Z > z) | Y] > x)] \quad (10)$$

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- Then, to discern how  $X$  and  $Y$  individually affect  $Z$ , we need to add a second parameter, say  $x$ , to extend the distribution to the meta distribution:

$$\bar{F}_{\llbracket Z|Y \rrbracket}(z, x) = \mathbb{E}[\mathbb{1}(\mathbb{E}[\mathbb{1}(Z > z) | Y] > x)] \quad (10)$$

or

$$\bar{F}_{\llbracket Z|Y \rrbracket}(z, x) = \mathbb{P}(\mathbb{P}(Z > z | Y) > x) \quad (11)$$

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- Hence the meta distribution (MD) is defined by first conditioning on part of the randomness.
  - It has two parameters, the distribution has one parameter, and the average has zero parameters.

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## Natural Progression and Back

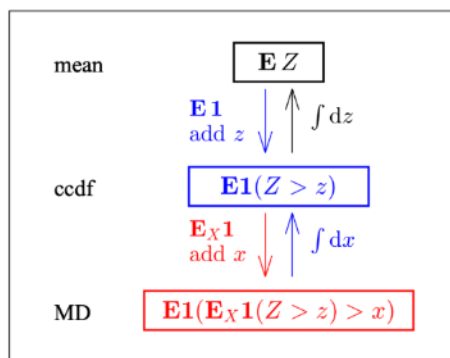


Figure: Going back and forth between mean, CCDF, and MD.

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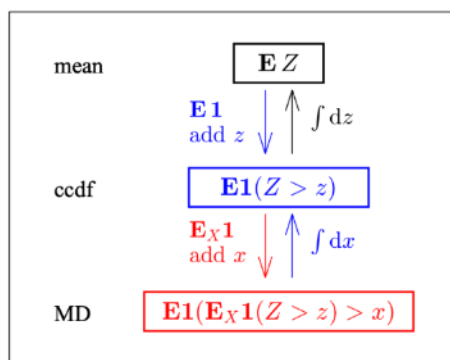


Figure: Going back and forth between mean, CCDF, and MD.

Let  $U = \mathbb{P}(Z > z|Y)$ . Then, the mean of  $U$  is the distribution of  $Z$  and the distribution of  $U$  is the MD.

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## Example

Let us consider

$$Z = \frac{X}{Y}$$

where  $X \sim \text{exponential}(\lambda_X)$  and  $Y \sim \text{exponential}(\lambda_Y)$ .

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$$1 - x^{\frac{\mu_Y}{z\mu_X}} \quad (12)$$

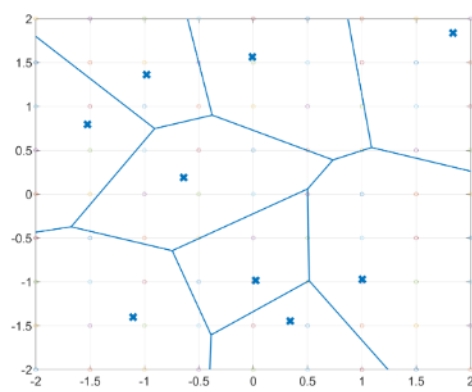
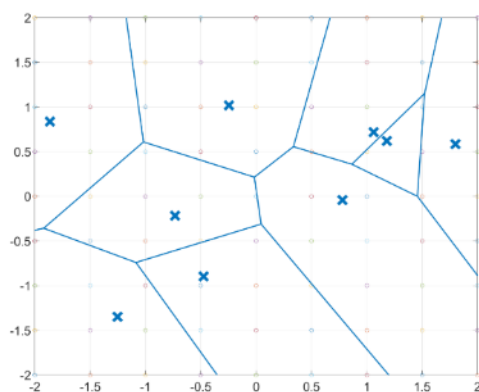
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## Example in a Wireless Network



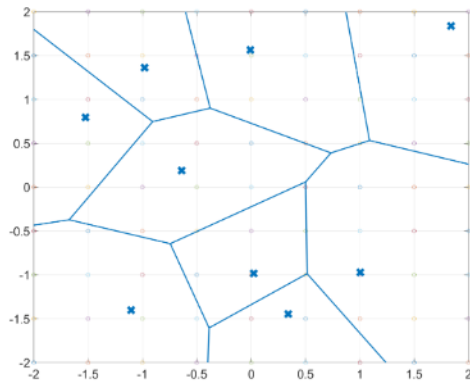
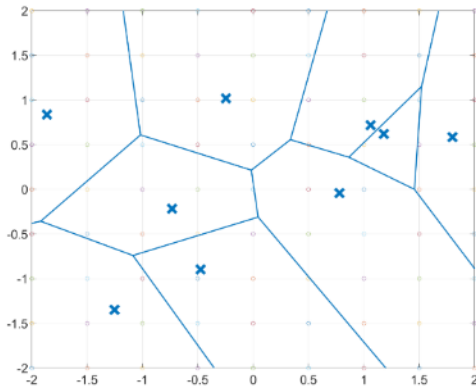
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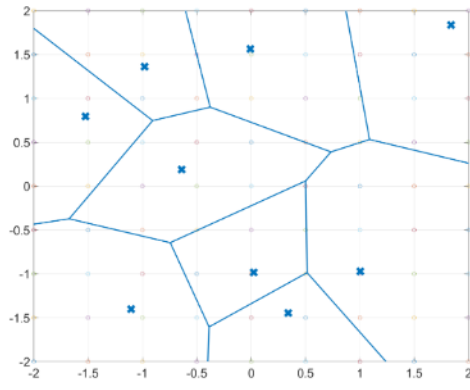
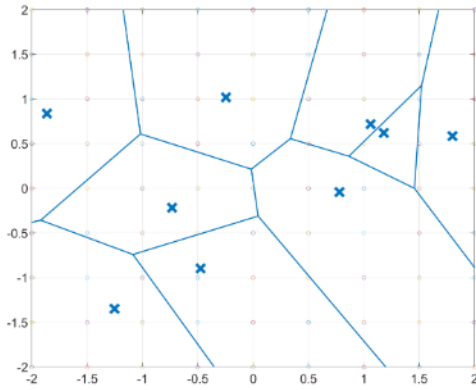
## Example in a Wireless Network



- PPP  $\implies$  nearest BS at distance at  $\sqrt{Y}$  where,  $Y \sim \text{exponential}(\pi\lambda)$ .

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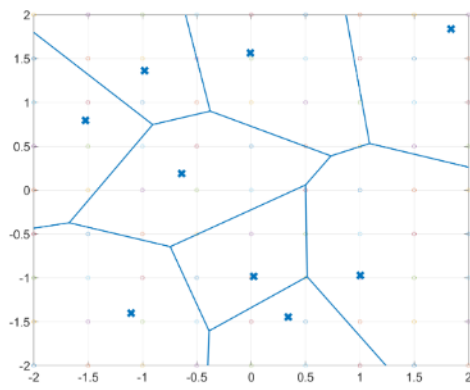
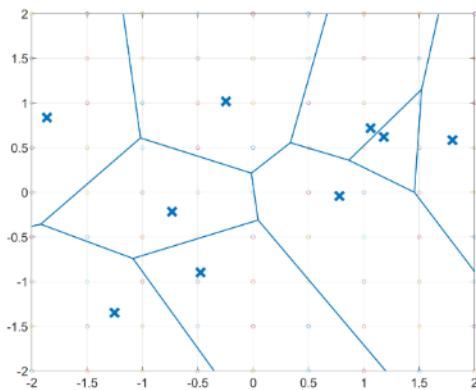
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## Example in a Wireless Network



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- Fading  $\implies$  Rx power  $\sim \text{exponential}(1)$ .
- Rx power  $Z = \frac{X}{Y}$ .

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## Example in a Wireless Network

Ergodicity  $\implies$  Let's talk about one realization (with  $\lambda = 1$ ) and ask  $\mathbb{P}(Z > 1|\Phi)$ ?

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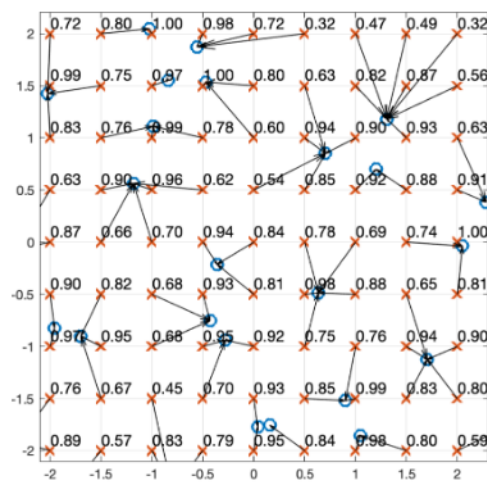
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## Example in a Wireless Network

- Histogram of all the user's probabilities?

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## Example in a Wireless Network

- Histogram of all the user's probabilities?
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- To know their distribution, we need to consult the MD.
- In contrast, without the MD, we have no information about the disparity between the users.
- Their personal probabilities could all be well concentrated around 0.76, or some could have probabilities near 0 and others near 1.
- Only the MD can reveal the performance of user percentiles, such as the "5% user" performance, which is the performance that 95% of the users achieve but 5% do not.

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## Let's return to BLCP - SINR Model

- $\Phi$  be a point process, the signal-to-interference-plus-noise ratio (SINR)  $\xi(r_0)$  is

$$\xi(r_0) = \frac{\xi_0 \|\mathbf{x}_1\|^{-\alpha} h_1}{1 + \xi_0 \sum_{\mathbf{x} \in \Phi \setminus \{\mathbf{x}_1\}} \|\mathbf{x}\|^{-\alpha} h_{\mathbf{x}}}, \quad (13)$$

$\xi_0$  is a constant that takes into account the transmit power, AWGN noise, path-loss constant, as well as the transmit and receive antenna gains.

For ease of notation, let us represent  $\|\mathbf{x}_i\|$  by  $d_x$ .

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Furthermore, let us assume ALOHA ( $p$ ).

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## Conditional Success Probability

The conditional success probability, i.e.,

$$P_s(\gamma) = \mathbb{P}(\xi(r_0) \geq \gamma | \Phi)$$

is a random variable due to the random  $\Phi$ .

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Its CCDF, called the meta distribution of the SINR, is given as

$$\mathcal{P}_M(\gamma, \beta) = \mathbb{P}(P_s(\gamma) \geq \beta) = \mathbb{P}(\mathbb{P}(\xi(r_0) \geq \gamma | \Phi) \geq \beta). \quad (14)$$

which is a function of two parameters  $\gamma \geq 0$  and  $0 \leq \beta \leq 1$ .

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## Conditional Success Probability

Let the set of locations of the interfering nodes be denoted by  $\mathcal{C} \subset \Phi'$ .

$$\begin{aligned}
 P_s(\gamma) &= \mathbb{P}(\xi(r_0) \geq \gamma \mid \Phi) = \mathbb{P}\left[\frac{\xi_0 d_1^{-\alpha} h_1}{1 + \xi_0 \sum_{\mathbf{x} \in \Phi'} h_{\mathbf{x}} d_{\mathbf{x}}^{-\alpha} \mathbf{1}(\mathbf{x} \in \mathcal{C})} \geq \gamma \mid \Phi\right] \\
 &= \mathbb{P}\left[h_1 > \frac{\gamma + \gamma \xi_0 \sum_{\mathbf{x} \in \Phi'} h_{\mathbf{x}} d_{\mathbf{x}}^{-\alpha} \mathbf{1}(\mathbf{x} \in \mathcal{C})}{\xi_0 d_1^{-\alpha}}\right] \stackrel{(a)}{=} \mathbb{E}_{h_{\mathbf{x}}}\left[\exp\left(\frac{-\gamma - \gamma \xi_0 \sum_{\mathbf{x} \in \Phi'} h_{\mathbf{x}} d_{\mathbf{x}}^{-\alpha} \mathbf{1}(\mathbf{x} \in \mathcal{C})}{\xi_0 d_1^{-\alpha}}\right)\right] \\
 &= \exp\left(\frac{-\gamma}{\xi_0 d_1^{-\alpha}}\right) \mathbb{E}_{h_{\mathbf{x}}}\left[\exp\left(\frac{-\gamma \xi_0 \sum_{\mathbf{x} \in \Phi'} h_{\mathbf{x}} d_{\mathbf{x}}^{-\alpha} \mathbf{1}(\mathbf{x} \in \mathcal{C})}{\xi_0 d_1^{-\alpha}}\right)\right] \\
 &= \exp\left(\frac{-\gamma}{\xi_0 d_1^{-\alpha}}\right) \left(\prod_{\mathbf{x} \in \Phi'} p \mathbb{E}_{h_{\mathbf{x}}}\exp\left(\frac{-\gamma \xi_0 d_{\mathbf{x}}^{-\alpha} h_{\mathbf{x}}}{\xi_0 d_1^{-\alpha}}\right) + 1 - p\right) \\
 &\stackrel{(b)}{=} \exp\left(\frac{-\gamma}{\xi_0 d_1^{-\alpha}}\right) \left(\prod_{\mathbf{x} \in \Phi'} \frac{p}{1 + \frac{\gamma d_{\mathbf{x}}^{-\alpha}}{d_1^{-\alpha}}} + 1 - p\right).
 \end{aligned}$$

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## Moments

- In general, directly deriving the distribution of the random variable  $P_s(\gamma)$  is most likely impossible.

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# Moments

- In general, directly deriving the distribution of the random variable  $P_s(\gamma)$  is most likely impossible.
- The standard approach to circumvent this challenge is by first deriving its moments and then transforming them to the distribution.
- Moments reveal key features:
  - First moment:  $M_1 = \mathbb{E}_\Phi [P_s(\gamma)] \rightarrow$  standard success probability.
  - $-1^{\text{th}}$  moment:  $M_{-1} = \mathbb{E}_\Phi \left[ \frac{1}{P_s(\gamma)} \right] \rightarrow$  mean local delay.
  - Similarly, variance etc.

## First Moment in a BLCP

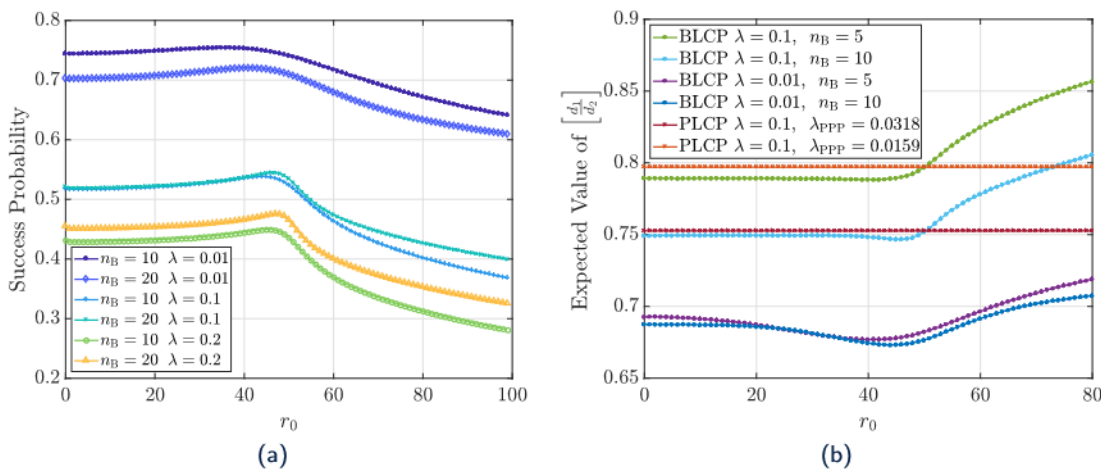


Figure: (a) Success probability with respect to  $r_0$ . Here  $R = 50$ . (b)  $\mathbb{E} \left[ \frac{d_1}{d_2} \right]$  with respect to  $r_0$ .

## $-1^{\text{th}}$ Moment

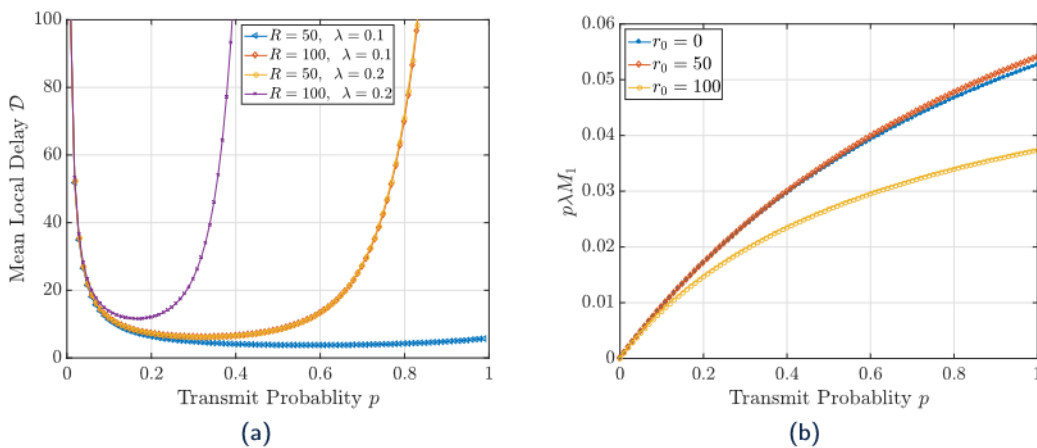


Figure: (a) Mean local delay with respect to the transmit probability for different values of  $R$  and  $\lambda$ . Here,  $r_0 = 0$ . (b) Successful transmission density.



## Optimal Access Probability

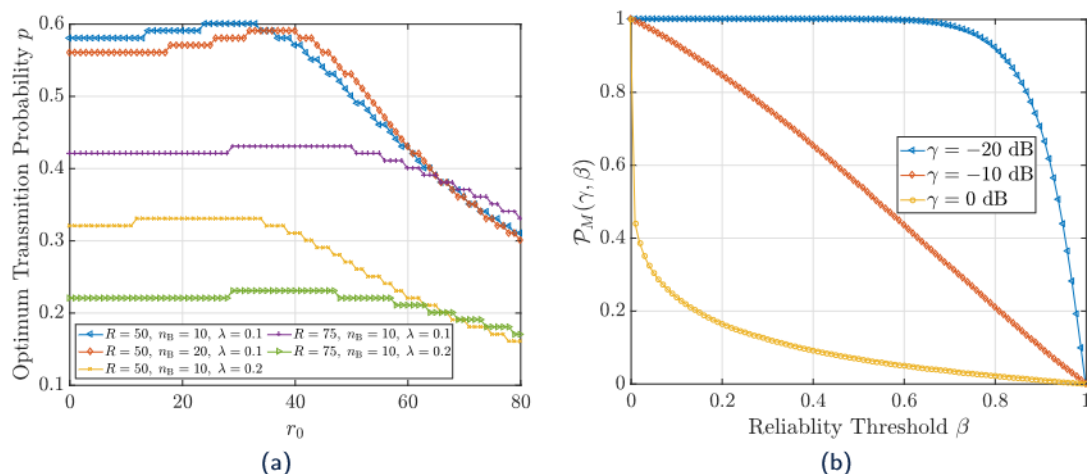


Figure: (a) Optimal transmit probability for minimizing the mean local delay. (b) SINR meta distribution.

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## Summary and Conclusions

- BLP and BLCP: New line process and Cox models that take into account the non-homogeneity of lines in a Euclidean plane.

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## Summary and Conclusions

- BLP and BLCP: New line process and Cox models that take into account the non-homogeneity of lines in a Euclidean plane.
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  - Rigorously studied in Part 2.

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- Inhomogeneity in the street network results in the adaptation of optimal wireless protocols - Tx probability, Load-balancing, automotive radar, etc.
  - Rigorously studied in Part 2.
- Questions:
  - How accurate is BLP/BLCP? - Working on it.
  - Nearest point in the  $L_1$  sense? Percolation questions.
  - How to integrate the BLP model with existing street models?

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## Thanks!

Questions?

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